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## Skip lists: A randomized dictionary

The exposition is based on the following sources, which are all recommended reading:

1. Pugh: Skip lists: a probabilistic alternative to balanced treed. Proceedings WADS, LNCS 382, 1989, pp. 437-449
2. Sedgewick: Algorithmen in C++, 2002, Pearsons, (Chapter 13.5)
3. Lecture Script from Michiel Smid, University of the Saarland.

## Introduction

We consider the so called dictionary problem. Given a set $S$ of ral numbers, store them in a data structure such that the following three opertions can be performed efficiently:

- Search( $\mathbf{x}$ ): Given the real number $x$, report the maximal element of $S \cup\{-\infty\}$ that is at most equal to $x$.
- Insert(x): Given a real number $x$ insert it into the data structure.
- Delete(x): Given a real number $x$ delete it into the data structure.


## Introduction

The standard data structures for this problem is the balanced binary tree. It supports all the above operations in worst case time $O(\log n)$ and uses $O(n)$ space.

Well known classes of balanced trees are for example AVL-trees, BB[ $\alpha]$-trees and red-black-trees. In order to maintain their worst case time behaviour all those data structures need more or less elaborate rebalancing operations which make an implementation non-trivial, which in turn leads not to the best practical run times.

## Introduction

We will introduce in this lecture an alternative, randomized data structure, the skip list. It uses in expectation linear space and supports the above dictionary operation in expected time $O(\log n)$ with high probability.

Why do we do this? We will see that the data structure is conceptually much simpler and more elegant than balanced trees. Nevertheless we will exchange a worst case runtime against an expected run time.

## Introduction

However, the analysis will show that skip lists behave very well and are very fast in practice (the difference is similar to the deterministic merge sort and the randomized quicksort algorithm).

The goal of this lecture is to a) introduce you to the data structure, b) show you how to analyze the randomized run time, and c) to introduce you to tail estimates using Chernoff bounds.

## Skip lists

Throughout the lecture we assume that we can generate random, independent bits in unit time. Let $S$ be a set of $n$ real numbers. Then we construct a sequence of sets $S_{1}, S_{2}, \ldots$ as follows:

1. For each element $x \in S$, flip a coin until zero comes up.
2. For each $i \geq 1, S_{i}$ is the set of elements in $S$ for which we flipped the coin at least $i$ times.

Let $h$ be the number of sets that are constructed. Then it is clear that

$$
\emptyset=S_{h} \subseteq S_{h-1} \subseteq S_{h-2} \subseteq \cdots \subseteq S_{2} \subseteq S_{1}=S
$$

## Skip lists

The skip list for $S$ consists then of the following:

1. For each $1 \leq i \leq h$, the elements of $S_{i} \cup\{-\infty\}$ are stored in a sorted linked list $L_{i}$.
2. For each $1<i \leq h$, there is a pointer from each $x \in L_{i}$ to its occurrence in $L_{i-1}$.

## Skip lists

Here is an example. Suppose $S=\{1,2,5,7,8,9,11,12,14,17,19,20\}$. Flipping coins might lead to $S_{1}=S, S_{2}=\{1,2,5,8,11,17,20\}, S_{3}=\{2,5,11,20\}, S_{4}=$ $\{11\}$, and $S_{5}=\emptyset$.


## Searching skip lists

We can now implement the search for $x$ as follows:

1. Let $y_{h}$ be the only element in $L_{h}$.
2. For $i=h, h-1, \ldots, 2$
(a) Follow the pointer from $y_{i}$ in $L_{i}$ to its occurence in $L_{i-1}$.
(b) Starting in $y_{i-1}$, walk to the right along $L_{i-1}$, until an element is reached that is larger than $x$ or the end of $L_{i-1}$ is reached. Let now $y_{i-1}$ be the last encountered element in $L_{i-1}$ that is at most equal to $x$.
3. Output $y_{1}$.

## Searching skip lists

The following figure illustrates the search step (we search for element 10):


It is not hard to imagine how the insertion and deletion operations work on a skip list.

## Inserting into a skip list

For inserting an element $x$ into the dictionary we proceed as follows:

1. Run the search algorithm for $x$. Let $y_{1}, y_{2}, \ldots, y_{h}$ be the elements of $L_{1}, L_{2}, \ldots, L_{h}$ that are computed while searching. If $x=y_{1}$, the $x \in S$ and nothing has to be done. Hence assume that $x \neq y_{1}$.
2. Flip a coin until a zero comes up. Let / be the number of coin flips.
3. For each $1 \leq i \leq \min (I, h)$, add $x$ to the list $L_{i}$ immediately after $y_{i}$.
4. If $I \geq h$, then create new lists $L_{h+1}, \ldots, L_{l+1}$ storing the sets $S_{h+1} \cup\{-\infty\}$, where each set contains $x$ except for $S_{I+1}$ which is empty.
5. For each $1<i \leq I$, give $x$ in $L_{i}$ a pointer to its occurrence in $L_{i-1}$.
6. If $I \geq h$, then for each $h+1 \leq i \leq I+1$, give $-\infty$ in $L_{i}$ a pointer to its occurrence in $L_{i-1}$.
7. Set $h=\max (h, l+q)$.


## Deleting from a skip list

1. Run the search algorithm for $x$. Let $y_{1}, y_{2}, \ldots, y_{h}$ be the elements of $L_{1}, L_{2}, \ldots, L_{h}$ that are computed while searching. If $x \neq y_{1}$, the $x \notin S$ and nothing has to be done. Hence assume that $x=y_{1}$.
2. For each $1 \leq i \leq h$ such that $x=y_{i}$ delete $y_{i}$ from the list $L_{i}$.
3. For $i=h, h-1, \ldots$ : if $L_{i-1}$ only stores $-\infty$, delete the list $L_{i}$ and set $h=h-1$.

## Why are skiplists efficient? The intuition

We have seen, that mostly what we do in skip lists is to search. The rebalancing is done by throwing a coin a few times and making local changes along the search path.

How expensive is the search? It is the sum over all traversed path length at each level. We expect there to be $\approx \log n$ levels. At each level we travel to the right. However for a fixed level we do not expect to do this long, since this would imply, that all the elements are not in the level above.

Hence we expect to spend a constant amount of time at each level which would add up to a total search time of $O(\log n)$. We will now prove this more formally.

## Why are skip lists efficient? The proofs

The size of a skip list and the running times of the search and update algorithms are random variables. We will prove that their expected values are bound by $O(n)$ and $O(\log n)$ respectively.

Recall that $h$ denotes the number of sets $S_{i}$ that result from our probabilistic construction. How can we derive and upper bound for $h$ ?

Let $x$ be an element of $S$ and $h(x)$ be the number of sets $S_{i}$ that contain $x$. Then $h(x)$ is a random variable distributed accoording to a geometric distribution with $p=1 / 2$. Hence $\operatorname{Pr}(H(x)=k)=(1 / 2)^{k}$ and $E(h(x))=2$. That means if we look at a specific element we only expect it to be in $S_{1}$ and $S_{2}$.

## Why are skip lists efficient? The proofs

Clearly $h=1+\max \{h(x): x \in S\}$. From $E(h(x))=2$ for any $x \in S$, however, we cannot conclude that the expected value of $h$ is three.

We can estimate $E(h)$ as follows. Again consider a fixed $x \in S$. It follows that for any $k \geq 1, h(x) \geq k$ if and only if the first $k-1$ coin flips produced a one. That is $\operatorname{Pr}\left(h(x) \geq k=(1 / 2)^{k-1}\right.$. In addition it is clear that $h \geq k+1$ if and only if there is a $x \in S$ such that $h(x) \geq k$. Hence

$$
\operatorname{Pr}(h \geq k+1) \leq n \cdot \operatorname{Pr}(h(x) \geq k)=\frac{n}{2^{k-1}}
$$

## Why are skip lists efficient? The proofs

This estimate does not make sense for $k<1+\log n$. For those values of $k$ we can use the trivial upper bound $\operatorname{Pr}(h \geq k+1) \leq 1$. Then $E(h)$ equals:

$$
\sum_{k=0}^{\infty} \operatorname{Pr}(h \geq k+1)=\sum_{k=0}^{\lceil\log n\rceil} \operatorname{Pr}(h \geq k+1)+\sum_{k=1+\lceil\log n\rceil}^{\infty} \operatorname{Pr}(h \geq k+1)
$$

(exercise: proof the first equality, that is $E(X)=\sum_{k=1}^{\infty} \operatorname{Pr}(X \geq k)$ for a random variable $X$ that takes values $\{0,1,2, \ldots\}$.)

## Why are skip lists efficient? The proofs

The first summation on the right hand side is at most $1+\lceil\log n\rceil$. The second sum can be bounded from above by:

$$
\sum_{k=1+\lceil\log n\rceil}^{\infty} \frac{n}{2^{k-1}}=n(1 / 2)^{\lceil\log n\rceil-1} \leq n(1 / 2)^{\log n-1} \leq 2
$$

Hence we have proved that $E(h) \leq 3+\lceil\log n\rceil$.

## Why are skip lists efficient? The proofs

The expected size of a skip list can easily be computed. Let $M$ denote the total size of the sets $S_{1}, S_{2}, \ldots, S_{h}$. Then $M=\sum_{x \in S} h(x)$ and by linearity of expectation:

$$
E(M)=\sum_{x \in S} E(h(x))=\sum_{x \in S} 2=2 n .
$$

If $M^{\prime}$ denotes the total number of nodes in a skip list, then $M^{\prime}$ is equal to $M$ plus $h$. Hence

$$
E\left(M^{\prime}\right)=E(M+h)=E(M)+E(h) \leq 2 n+3+\lceil\log n\rceil .
$$

## Why are skip lists efficient? The proofs

What is left to do is to estimate the search costs.
Let $x$ be a real number and let $C_{i}$ denote the number of elements in the list $L_{i}$ that are inspected when searching for $x$ (We do not count the element of $L_{i}$ at which the algorithm starts walking to the right. Hence, $C_{i}$ counts comparisons between $x$ and elements of $S$.) The search cost is then proportional to $\sum_{i=1}^{h}\left(1+C_{i}\right)$.

Again we cannot use linearity of expectation since $h$ is a random variable. Again the trick is to fix an integer $A$ and analyze the search cost up to a level $A$ and above level $A$ separately (and differently).

## Why are skip lists efficient? The proofs

We first estimate the search level above $A$, i.e., the total costs in the lists $L_{A+1}, L_{A+2}, \ldots, L_{h}$. Since the cost is at most equal to the total size of these lists, its expected value is at most equal to the expected value of $M_{A}:=\sum_{i=A+1}^{h}\left|L_{i}\right|$.

How do we estimate this value? We first note that the lists $L_{i}, A+1 \leq i \leq h$, form a skip list for $S_{A+1}$. Hence we have:

$$
E\left(M_{A}\right)=\sum_{k=0}^{n} E\left(M_{A}| | S_{A+1} \mid=k\right) \cdot \operatorname{Pr}\left(\left|S_{A+1}\right|=k\right)
$$

where $E\left(M_{A}| | S_{A+1} \mid=k\right)$ is the expected size of a skip list with $k$ elements. We have already seen that this is $O(k)$.

## Why are skip lists efficient? The proofs

Hence we only need to compute $\operatorname{Pr}\left(\left|S_{A+1}\right|=k\right)$. Since $\left|S_{A+1}\right|=k$ if and only if out of the $n$ elements of $S$ exactly $k$ reach the level $A+1$, we have:

$$
\operatorname{Pr}\left(\left|S_{A+1}\right|=k\right)=\binom{n}{k}\left(\frac{1}{2}\right)^{A k}\left(1-\left(\frac{1}{2}\right)^{A}\right)^{n-k} .
$$

Setting $p=\frac{1}{2}^{A}$, we infer that the expected value of $M_{A}$ is proportional to:

$$
\begin{aligned}
\sum_{k=0}^{n} k \cdot\binom{n}{k} p^{k}(1-p)^{n-k} & =\sum_{k=1}^{n} n \cdot\binom{n-1}{k-1} p^{k}(1-p)^{n-k} \\
& =n \cdot p \sum_{k=0}^{n-1}\binom{n-1}{k} p^{k}(1-p)^{n-1-k} \\
& =n \cdot p(p+(1-p))^{n-1} \\
& =n \cdot p
\end{aligned}
$$

Hence the expected search cost above level $A$ is bounded by $O\left(n / 2^{A}\right)$.

## Why are skip lists efficient? The proofs

Next we estimate the expected search cost in the lists $L_{1}, L_{2}, \ldots, L_{A}$. Recall that $C_{i}$ is the number of elements searched when searching for $x$. We use again conditional expectation. Let $l_{i}(x)$ be the number of elements in $L_{i}$ that are at most equal to $x$. Then

$$
E\left(C_{i}\right)=\sum_{k=1}^{n} E\left(C_{i} \mid I_{i}(x)=k\right) \cdot \operatorname{Pr}\left(l_{i}(x)=k\right) .
$$

## Why are skip lists efficient? The proofs

Assume that $l_{i}(k)=k$. Also assume that there is an element in $L_{i}$ that is larger than $x$. Then $C_{i}=j$ if and only if the largest $j-1$ elements of $L_{i}$ that are at most equal to $x$ do not appear in $L_{i+1}$, but the element that immediately precedes these $j-1$ elements does appear in $L_{i+1}$.

## Why are skip lists efficient? The proofs

Hence

$$
\operatorname{Pr}\left(C_{i}=j \mid l_{i}(x)=k\right) \leq\left(\frac{1}{2}\right)^{j-1}, 0 \leq j \leq k
$$

This inequality also holds if $x$ is at least equal to the maximal element of $L_{i}$. From this we obtain:

$$
\begin{aligned}
E\left(C_{i} \mid l_{i}(x)=k\right) & =\sum_{j=0}^{k} j \cdot \operatorname{Pr}\left(C_{i}=j \mid I_{i}(x)=k\right) \\
& \leq \sum_{j=0}^{k} \frac{j}{j^{j-1}} \\
& \leq 4 .
\end{aligned}
$$

(exercise.)

## Why are skip lists efficient? The proofs

This, in turn implies that

$$
E\left(C_{i}\right) \leq \sum_{k=1}^{n} 4 \cdot \operatorname{Pr}\left(l_{i}(x)=k\right)=4
$$

It follows that the expected search cost up to level $A$ is proportional to:

$$
E\left(\sum_{i=1}^{A}\left(1+C_{i}\right)\right)=\sum_{k=1}^{A}\left(1+E\left(C_{i}\right)\right) \leq 5 A
$$

Summarizing we have shown that the expected search time for element $x$ is bounded by:

$$
O\left(\frac{n}{2^{A}}+A\right) .
$$

Setting $A$ to $\log n$ we obtain the required bound of $O(\log n)$.

## Tail estimates: Chernoff bounds

So far we proved bounds on the expected size, search time and update time for a skip list. In this section we conder so called tail estimates.

That is, we estimate the probability that the actual search time deviates significantly from its expected value. For example assume for a moment that the constant in the $O(\log n)$ term for the search time is one. Then we want to estimate the probability that the actual search time is at least $t \cdot \log n$.

We could derive an estimate using Markov's inequality.

## Tail estimates: Chernoff bounds

Lemma 1. Let $X$ be a random variable that takes non-negative values, and let $\mu$ be the expected value of $X$. Then for any $t>0, \operatorname{Pr}(Y \geq t \mu) \leq \frac{1}{t}$.
Proof: Let $s=t \mu$. Then

$$
\begin{aligned}
\mu & =\sum_{x} x \cdot \operatorname{Pr}(X=x) \\
& \geq \sum_{x \geq s} x \cdot \operatorname{Pr}(X=x) \\
& \geq \sum_{x \geq s} s \cdot \operatorname{Pr}(X=x) \\
& =s \cdot \operatorname{Pr}(X \geq s)
\end{aligned}
$$

## Tail estimates: Chernoff bounds

Hence the probability that the actual search time is at least $t \cdot \log n$ is less than or equal to $1 / t$.

This is not very impressive. The probablity that the search time is more than 100 times its expected value is at most $1 / 100$. So if this bounds was tight one search in a hundred takes more than 100 times the time of the average search.

## Tail estimates: Chernoff bounds

In this section we will se that Chernoff bounds give a much tighter estimate. We will prove that the probablity that the search time exceeds $t \cdot \log n$ is less than or equal to $\approx n^{-t / 8}$ for $t \geq 5$.

Hence in a skip list of 1000 elements, the probability that the search time is more than 100 times its expected value is $10^{-38}$ which in practice means, it will never occur. (Even for $t=50$ the bound is still $10^{-19}$, and for $t=10$ the probability is still only $\approx 2 \cdot 10^{-4}$ ).

## Tail estimates: Chernoff bounds

Markov's inequality holds for any non-negative random variable. The Chernoff technique applies to random variables $X$ that can be written as the sum $\sum_{i=1}^{n} X_{i}$ of mutually independent random variables $X_{i}$.
(Variables are called (mutually) independent if their joint density function is the product of the individual density functions. Beware that mutual independence is different than pairwise independence! (exercise)).

In such cases much better bounds can be obtained.

## Tail estimates: Chernoff bounds

So let $X_{1}, X_{2}, X_{3} \ldots, X_{n}$ be a sequence of mutually independent random variables and let $X=\sum_{i=1}^{n} X_{i}$.

The moment generating function (mgf) for a (discrete) random variable $Y$ is defined as

$$
m_{Y}(\lambda)=E\left(e^{\lambda Y}\right)=\sum_{y} e^{\lambda y} \cdot \operatorname{Pr}(Y=y)
$$

As the name suggests the function is used to easily generate the moments of the random variable $Y$. Clearly $m(0)=1$ and it is easy to show that $\mu=m^{\prime}(0)$ and $\sigma^{2}=m^{\prime \prime}(0)-\mu^{2}$ (exercise).

## Tail estimates: Chernoff bounds

In the case of $X$, which is a sum of $n$ independent variables, $m_{X}(\lambda)=\prod_{i=1}^{n} m_{X_{i}}(\lambda)$ and of course the mean value of $X$ is the derivate of the mgf at position 0 which is simply the product of all means of the $X_{i}$.

Or written down:

$$
E\left(e^{\lambda X}\right)=E\left(e^{\lambda\left(X_{1}+\cdots+X_{n}\right)}\right)=\prod_{i=1}^{n} E\left(e^{\lambda X_{i}}\right) .
$$

## Tail estimates: Chernoff bounds

Now let $s>0$ and $\lambda>0$. Since $X \geq s$ if and only if $e^{\lambda X} \geq e^{\lambda s}$, we have $\operatorname{Pr}(X \geq s)=\operatorname{Pr}\left(e^{\lambda X} \geq e^{\lambda s}\right)$. By applying Markov's inequality to the non negative random variable $e^{\lambda X}$, we get

$$
\operatorname{Pr}(X \geq s)=\operatorname{Pr}\left(e^{\lambda X} \geq e^{\lambda s}\right) \leq e^{-\lambda s} \cdot E\left(e^{\lambda X}\right)
$$

This yields:

$$
\operatorname{Pr}(X \geq s) \leq e^{-\lambda s} \cdot \prod_{i=1}^{n} E\left(e^{\lambda X_{i}}\right), \text { for } s>0 \text { and } \lambda>0
$$

This is the basic inequality we work with. To estimate $\operatorname{Pr}(X \geq s)$ we need bound on $E\left(e^{\lambda X_{i}}\right)$. Of course those bounds depend on the probability distribution of $X_{i}$. We will now illustrate the technique using the geometric distribution with parameter $p=1 / 2$.

## Tail estimates: Chernoff bounds

Let $T$ be the number of flips we need until a one comes up in a series of coin flips. Then $\operatorname{Pr}(T=k)=(1 / 2)^{k}$ for $k \geq 1$ and $E(T)=2$. Now assume we are interested in $T_{n}$ which is the number of flips we need until we obtain a one exactly $n$ times (i.e. $\left.T=T_{1}\right)$.

## Tail estimates: Chernoff bounds

If we define the random variables $X_{i}$ as the number of flips between the ( $i-1$ )-st (excluding) and the $i$-th one (including), then $X_{i}$ is distributed according to a geometric distribution. (This property is also called the memoryless property of the geometric or exponential distribution).

Then $T_{n}=\sum_{i=1}^{n} X_{i}$, where each $X_{i}$ is distributed according to a geometric distribution and the expected value is $E\left(T_{n}\right)=2 n$, and Markov's inequality gives $\operatorname{Pr}\left(T_{n} \geq(2+t) n \leq \frac{2}{2+t}\right)$.

## Tail estimates: Chernoff bounds

For $0<\lambda<\log 2$ we have

$$
E\left(e^{\lambda X_{i}}\right)=\sum_{k=1}^{\infty} e^{\lambda k} \cdot \operatorname{Pr}\left(X_{i}=k\right)=\sum_{k=1}^{\infty}\left(e^{\lambda} / 2\right)^{k}=\frac{e^{\lambda}}{2-e^{\lambda}}
$$

We now apply our basic inequality with $s=(2+t) n$, where $t>0$ and get

$$
\operatorname{Pr}\left(T_{n} \geq(2+t) n\right) \leq e^{-\lambda(2+t) n}\left(\frac{e^{\lambda}}{2-e^{\lambda}}\right)^{n}=\left(\frac{e^{-\lambda(1+t)}}{2-e^{\lambda}}\right)^{n}
$$

Now we choose $\lambda$ such that the term on the right hand side is minimized (exercise) and find $\lambda=\log \left(1+\frac{t}{2+t}\right)$.

## Tail estimates: Chernoff bounds

Hence we have

$$
\operatorname{Pr}\left(T_{n} \geq(2+t) n\right) \leq(1+t / 2)^{n}\left(1-\frac{t}{2+2 t}\right)^{(1+t) n} .
$$

Since $1-x \leq e^{-x}$ for all $x$, we have

$$
\left(1-\frac{t}{2+2 t}\right)^{1+t} \leq\left(e^{\left.\frac{-t}{2+2 t}\right)^{1+t}}=e^{-t / 2} .\right.
$$

Moreover, $1+t / 2 \leq e^{t / 4}$ for $t \geq 3$. This proves that for $t \geq 3$

$$
\operatorname{Pr}\left(T_{n} \geq(2+t) n\right) \leq e^{t n / 4} \cdot e^{-t n / 2}=e^{-t n / 4}
$$

Compare this with the bound obtained from Markov's inequality (which was $\frac{2}{2+t}$ )!

## Tail estimates: Chernoff bounds

We can subsume our finding in the following theorem:
Theorem 2. Let $X_{1}, X_{2}, \ldots, X_{n}$ be mutually indpendent random variables and assume that each $X_{i}$ is distributed according to a geometric distribution. Let $T_{n}=\sum X_{i}$, then $E\left(T_{n}\right)=2 n$ and for any $t \geq 3$ holds:

$$
\operatorname{Pr}\left(T_{n} \geq(2+t) n\right) \leq e^{t n / 4} \cdot e^{-t n / 2}=e^{-t n / 4}
$$

Corollary 3. Let $c \geq 1$ be a constant and let $m$ be a positive integer. Further let $n=c \cdot \ln m$. Then for any $s \geq 5$ it holds

$$
\operatorname{Pr}\left(T_{\lceil n\rceil} \geq s n\right) \leq m^{\frac{-(s-2) c}{4}}
$$

## Tail estimates: Skip lists

We can now use the above results to prove tail estimates for the size, search time and update time for a skip list. Consider again a skip list for $n$ elements. Let $M$ denote the total size of the sets $S_{i}, \ldots, S_{h}$ and let $M^{\prime}$ denote the total number of nodes of the skip list. Then $M^{\prime}=M+h$. We have seen that $E\left(M^{\prime}\right)=E(M+h) \leq$ $2 n+3+\lceil\log n\rceil$.

Now we are interested in the probability that $M^{\prime}$ is at least equal to $(2+t) n$. Clearly if $M^{\prime} \geq(2+t) n$, then $h \geq \frac{t n}{2}$ or $M \geq\left(2+\frac{t}{2}\right) n$. As a result,

$$
\operatorname{Pr}\left(M^{\prime} \geq(2+t) n\right) \leq \operatorname{Pr}\left(h \geq \frac{t n}{2}\right)+\operatorname{Pr}\left(M \geq\left(2+\frac{t}{2}\right) n\right)
$$

## Tail estimates: Skip lists

In the beginning of the lecture we estimated already the first probability on the right hand side and have:

$$
\operatorname{Pr}\left(h \geq \frac{t n}{2}\right) \leq \frac{n}{2^{\frac{t n}{2}-2}}=e^{\ln n+2 \ln 2-\frac{t n}{2} \ln 2} \leq e^{\frac{-t n}{8}}
$$

It remains to bound $\operatorname{Pr}\left(M \geq\left(2+\frac{t}{2}\right) n\right)$.
Remember that $h(x)$ was defined as the number of sets $S_{j}$ that contain $x$. Then the random variables $h(x)$ for $x \in S$ are all mutually independent and have a geometric distribution. Since $M=\sum_{x \in S} h(x)$ we can apply above theorem for geometric distributions and get:

$$
\operatorname{Pr}\left(M \geq\left(2+\frac{t}{2}\right) n\right) \leq e^{\frac{-t n}{8}} .
$$

That means it is extremly unlikely that a skip list deviates much from its expected size.

## Tail estimates: Skip lists

The analysis of the search time is a bit more complicated. Let $x \in \mathbb{R}$ and let $c \geq 1$ be a constant. We want to estimate the probability that the searching for $x$ takes more then $c \cdot \log n$ steps. We analyze this by considering the cost up to level $\lceil c \cdot \log n\rceil$ and above level $\lceil c \cdot \log n\rceil$ separately.

Let $A$ denote the number of nodes traversed in levels $1+\lceil c \cdot \log n\rceil, 2+\lceil c \cdot \log n\rceil, \ldots, h$ when searching for $x$. We proved before that the expected value of $A$ is bounded by $O\left(\frac{n}{2^{[c \cdot \log n\rceil}}\right)=O\left(\frac{1}{n^{c \mid n} 2-1}\right)$. Hence Markov's inequality yields:

$$
\operatorname{Pr}(A \geq 1) \leq E(A)=O\left(\frac{1}{n^{c \ln 2-1}}\right) .
$$

## Tail estimates: Skip lists

Now let $B$ be the number of nodes traversed in the lower levels. We want to express $B$ now as the sum of mutually independent random variables each distributed according to a geometric distribution.

In order to do this we give an slightly different but equivalent construction of a skip list.

Let $S_{1}=S=\left\langle s_{11}<s_{12}<\ldots<s_{1 n_{1}}\right\rangle$, where $n_{1}=n$. Moreover, let $l_{1}(x)$ be the number of elements of $S_{1}$ that are at most equal to $x$.

Let $i \geq 1$ and assume that we have constructed $S_{i}=\left\langle s_{11}<s_{12}<\ldots<s_{1 n_{i}}\right\rangle$ with $n_{i}=\left|S_{i}\right|$. Let $l_{i}(x)=\left|\left\{y \in S_{i}: y \leq x\right\}\right|$.

## Tail estimates: Skip lists

We do the following:

1. We flip our coin $n_{i}-l_{i}(x)$ times.
2. We flip the coin $I_{i}(x)$ times.
3. If the previous $l_{i}(x)$ flips only produced zeros, then we flip the coin until we get a one.

Let $f_{i}$ be the total number of flips made during these stages, and denote the outcome by $F_{i 1}, F_{i 2}, \ldots, F_{i f_{i}}$. Note that $f_{i} \geq n_{i}$. We proceed as follows:

## Tail estimates: Skip lists

1. Define $S_{i+1}=\left\{s_{i j}: 1 \leq j \leq n_{i}\right.$ and $\left.F_{i j}=1\right\}, n_{i+1}=\left|S_{i+1}\right|$ and $l_{i+1}(x)=\{y \in$ $\left.S_{i+1}: y \leq x\right\}$.
2. We define the random variable $X_{i}$ as the number of flips after stage 1 until the first 1 comes up.
3. We define the random variable $C_{i}$ as the minimum of $1+l_{i}(x)$ and $X_{i}$.

The construction stops as soon as the empty set $S_{i+1}$ has been constructed.

## Tail estimates: Skip lists

Hence we have now

$$
\emptyset=S_{h} \subset S_{h-1} \subseteq S_{h-2} \subseteq \ldots \subseteq S_{2} \subseteq S_{1}=S
$$

and random variables $X_{1}, X_{2}, \ldots, X_{h-1}$ and $C_{1}, C_{2}, \ldots, C_{h-1}$. This construction defines a probability distribution on skip lists, which is the same as that of our previous two constructions.

## Tail estimates: Skip lists

Lets go back to the analysis of the search time. For convenience, we define $C_{i}=0$ fo all $i \geq h$. Recall that $B$ is the number of nodes traversed in the levels $1,2, \ldots,\lceil c$. $\log n\rceil$. Hence we have $B=\sum_{i=1}^{[c \cdot \log n\rceil} C_{i}$. The means we have written $B$ as the sum of random variables.

However, since $C_{i}$ depends on $l_{i}(x)$, which depends on $n_{i}$ which in turn depends on $n_{i-1}$, we see that the variables $C_{1}, C_{2}, \ldots, C_{h-1}$ are not mutually independent.

## Tail estimates: Skip lists

On the other hand, it is easy to see that $C_{i} \leq X_{i}$ for all $1 \leq i \leq h-1$. Moreover, the random variables $X_{i}, 1 \leq i \leq h-1$ are mutually independent and each one is distributed according to a geometric distribution. For $i \geq h$, let $X_{i}$ also denote a random variable distributed according to a geometric distribution. Since

$$
B=\sum_{i=1}^{\lceil c \cdot \log n\rceil} C_{i} \leq \sum_{i=1}^{\lceil c \cdot \log n\rceil} x_{i},
$$

we infer

$$
P(B \geq s \cdot c \cdot \log n) \leq \operatorname{Pr}\left(\sum_{i=1}^{\lceil c \cdot \log n\rceil} X_{i} \geq s \cdot c \cdot \log n\right)
$$

The above Corollary immediately gives:

$$
P(B \geq s \cdot c \cdot \log n) \leq n^{\frac{-(s-2) c}{4}} \leq \frac{1}{n^{\frac{s c}{8}}},
$$

for $c \geq 1$ and $s \geq 5$.

## Tail estimates: Skip lists

Now we can give the tail estimate for the search time. Let $T$ denote the total number of nodes trversed when searching for $x$. Then $T=A+B$. Moreover, if $T \geq 1+5 c$. $\log n$, then $A \geq 1$ or $B \geq 5 x \cdot \log n$. Hence

$$
\operatorname{Pr}(T \geq 1+5 c \cdot \log n) \leq \operatorname{Pr}(A \geq 1)+\operatorname{Pr}(B \geq 5 c \cdot \log n) .
$$

Our results for the two probabilities on the right hand side imply for $c \geq 1$,

$$
\operatorname{Pr}(T \geq 1+5 c \cdot \log n)=O\left(n^{1-c \cdot \ln 2}+n^{-5 c / 8}\right) .
$$

Taking $c=t / 5$, where $t \geq 5$ we get

$$
\operatorname{Pr}(T \geq 1+t \log n)=O\left(n^{1-(t \cdot \ln 2) / 5}+n^{-t / 8}\right)
$$

